Maximising likelihood for multivariate Gaussian distributions

When you find the mean of some data, what you are really doing is finding a parameter which *maximises the likelihood*.

For instance, assume some points $\{\vec{x}_1, ..., \vec{x}_N\} \subset \mathbb{R}^d$ are normally distributed. The conditional probability of the data is

$$P(\vec{x}_i \mid \vec{\mu}, \Sigma) = \prod_{i=1}^N \frac{1}{\sqrt{\tau^d \det \Sigma}} \exp\left(-\frac{1}{2} (\vec{x}_i - \vec{\mu})^T \Sigma^{-1} (\vec{x}_i - \vec{\mu})\right)$$

which is also the *likelihood* of the parameters $\vec{\mu}$ and Σ . It is often easiler to manipulate the logarithm of the likelihood:

$$\log P = -\frac{1}{2} \log \left(\tau^d \det \Sigma\right) - \frac{1}{2} \sum_{i=1}^N \left(\vec{x}_i - \vec{\mu}\right)^T \Sigma^{-1} (\vec{x}_i - \vec{\mu})$$

Fitting $\vec{\mu}$ to data

To find the mean $\vec{\mu}$ which maximises the likelihood, note that log *P* is quadratic in $\vec{\mu}$, so the maximum occurs at the unique point where its derivative vanishes.

Consider the differential $\delta \log P$ induced by $\vec{\mu} \rightarrow \vec{\mu} + \delta \vec{\mu}$:

$$\begin{split} \delta \log P &= \frac{1}{2} \sum_{i=1}^{N} \left[(\delta \vec{\mu})^T \Sigma^{-1} (\vec{x}_i - \vec{\mu}) + (\vec{x}_i - \vec{\mu})^T \Sigma^{-1} \delta \vec{\mu} \right] \\ &= \sum_{i=1}^{N} \left[(\vec{x}_i - \vec{\mu})^T \Sigma^{-1} \right] \delta \vec{\mu} \\ &= \left[\sum_{i=1}^{N} \vec{x}_i - N \vec{\mu} \right]^T \Sigma^{-1} \delta \vec{\mu} \end{split}$$

If the likelihood is at a local maximum, then $\delta \log P$ must vanish for any $\delta \vec{\mu}$. This holds when:

$$\vec{\mu} = \frac{1}{N}\sum_{i=1}^N \vec{x}_i$$

Fitting Σ to data

To find the covariance matrix Σ which maximises the likelihood, consider the differential likelihood induced by $\Sigma \rightarrow \Sigma + \delta \Sigma$.

$$\delta \log P = -\frac{N}{2} \frac{\delta \det \Sigma}{\det \Sigma} - \frac{1}{2} \sum_{i=1}^{N} \left(\vec{x}_i - \vec{\mu} \right)^T \delta(\Sigma^{-1}) (\vec{x}_i - \vec{\mu})$$

Use the identities

$$\begin{split} &\delta \det A = \mathrm{tr} [A^{-1} \delta A] \det A \\ &\delta (\Sigma^{-1}) = \Sigma^{-2} \delta \Sigma = \delta \Sigma \, \Sigma^{-2} \end{split}$$

and take the trace to obtain

$$\mathrm{tr}[\delta \log P] = -\frac{N}{2} \operatorname{tr}[\Sigma^{-1} \delta \Sigma] - \frac{1}{2} \sum_{i=1}^{N} \mathrm{tr}[(\vec{x} - \vec{\mu})(\vec{x} - \vec{\mu})^T \Sigma^{-2} \delta \Sigma]$$

where we use the cyclic property of the trace in the last term.

If the likelihood is at a local maximum, then it vanishes for any $\delta\Sigma$. Since $\delta\Sigma$ is arbitrary, this scalar equality between trace implies equality between the matrices themselves:

$$-2\delta \log P = -N\Sigma^{-1}\delta\Sigma + \sum_{i=1}^{N} (\vec{x} - \vec{\mu})(\vec{x} - \vec{\mu})^{T}\Sigma^{-2}\delta\Sigma$$

This vanishes when the covariance matrix is given by:

$$\Sigma = \frac{1}{N}\sum_{i=1}^N (\vec{x}-\vec{\mu})(\vec{x}-\vec{\mu})^T$$